On the role of total variation in compressed sensing Exact reconstructions from highly incomplete Fourier data

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Sampling the Fourier transform

In many applications, we are required to recover some function $f \in L^2(\Omega)$, $\Omega \subset \mathbb{R}^d$, from pointwise evaluations of its Fourier transform:

$$\mathcal{F}f(\omega) = \int_{x\in\Omega} f(x)e^{-2\pi i\langle\omega,x\rangle} \mathrm{d}x.$$

- ▶ Medical imaging: Magnetic Resonance Imaging (MRI)/ Computed Tomography
- Astronomy: Radio Interferometry
- ▶ Biology: Electron/Fluorescence Microscopy

The Shannon Nyquist Sampling Theorem Whittaker 1929, Kotelnikov 1933, Shannon 1949

If f has support included in [-T, T], then for $\epsilon^{-1} \ge 2T$,

 $f = \epsilon \sum \mathcal{F} f(\epsilon n) e^{2\pi i \epsilon n}$, with L^2 convergence, $n \in \mathbb{Z}^2$

$$\mathcal{F}f = \sum_{n \in \mathbb{Z}^2} \mathcal{F}f(\epsilon n) \operatorname{sinc}\left(\frac{\cdot + n\epsilon}{\epsilon}\right)$$

with L^{∞}, L^2 convergence.

The Shannon Nyquist Sampling Theorem Whittaker 1929, Kotelnikov 1933, Shannon 1949

If f has support included in [-T, T], then for $\epsilon^{-1} \ge 2T$,

$$\begin{split} f &= \epsilon \sum_{n \in \mathbb{Z}^2} \mathcal{F}f(\epsilon n) e^{2\pi i \epsilon n \cdot}, \qquad \text{with } L^2 \text{ convergence,} \\ \mathcal{F}f &= \sum_{n \in \mathbb{Z}^2} \mathcal{F}f(\epsilon n) \text{sinc}\left(\frac{\cdot + n\epsilon}{\epsilon}\right), \qquad \text{with } L^\infty, L^2 \text{ convergence.} \end{split}$$

Under no further assumptions, the sampling rate ϵ^{-1} must be at least 2T, the Nyquist rate.

Sparsity

In the last few decades, sparsity has played a prominent role in image processing.



Total variation (1992), wavelets (1988), contour lets (2005), curvelets (2000), shearlets (2006), \dots

An intriguing experiment Candès, Romberg & Tao, 2006

with



 $\min_{z \in \mathbb{C}^{N \times N}} \left\| z \right\|_{TV} \text{ subject to } P_{\Omega} U z = P_{\Omega} U x$

►
$$Uz = \left(\sum_{j_1=1}^{N} \sum_{j_2=1}^{N} z_{j_1,j_2} e^{i2\pi(k_1j_1+k_2j_2)}\right)_{k_1,k_2=-\lfloor N/2 \rfloor,...,\lceil N/2 \rceil-1}$$
.
► $Dz = D_1z + iD_2z$ where

$$D_1 z = (x_{k+1,j} - z_{k,j})_{k,j=1}^N, \quad D_2 z = (z_{k,j+1} - z_{k,j})_{k,j=1}^N,$$
$$z_{N+1,j} := z_{1,j} \text{ and } z_{k,N+1} := z_{k,1}. \text{ Let } \|z\|_{TV} := \|Dz\|_1.$$

A theoretical explanation towards sub-Nyquist sampling

Candès, Romberg & Tao, 2006: Let $x \in \mathbb{C}^N$ be *s*-sparse in its discrete gradient and suppose we observe its 0^{th} Fourier coefficient plus $\mathcal{O}(s \log N)$ of its Fourier coefficients chosen uniformly at random. Then, with overwhelming probability, x is the unique solution to

 $\min_{z \in \mathbb{C}^N} \|z\|_{TV} \text{ subject to } P_{\Omega}Uz = P_{\Omega}Ux.$





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• $\mathcal{O}(s \log N)$ represents a substantial saving in the number of samples.

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▶ $\mathcal{O}(s \log N)$ represents a substantial saving in the number of samples.

▶ The sampling cardinality of $\mathcal{O}(s \log N)$ is optimal for *s*-sparse vectors.

Compressed sensing*

How can we recover an s-sparse vector $x \in \mathbb{C}^N$ from Uxwhere $U \in \mathbb{C}^{m \times N}$ and $m = \mathcal{O}(s \log N) \ll N$?

- Thousands of papers developing algorithms and random sensing matrices.
- But many applications where compressed sensing is of practical interest are constrained to Fourier sampling, and one of the most widely used sparsifying transform is the gradient operator.
 - Lustig et al. (2007) on MRI, Wiaux et al. (2009) on radio interferometry, Leary et al. (2013) on electron microscopy,

*Introduced in 2006 independently by Donoho and Candès, Romberg & Tao.

This talk:

In practice, signals are only approximately sparse and measurements are noisy. Given $y = P_{\Omega}Ux + \eta \in \mathbb{C}^m$ with $\|\eta\|_2 \leq \delta\sqrt{m}$, we seek to solve

 $\min_{z} \|z\|_{TV} \text{ subject to } \|P_{\Omega}Uz - y\|_2 \le \delta\sqrt{m}.$



Reconstructions from sampling 10% of the Fourier coefficients.

This talk: 2 key questions

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Reconstructions from sampling 10% of the Fourier coefficients.

- 1. What can we say about robustness to noise and stability to inexact sparsity under uniform random sampling at $\mathcal{O}(s \log N)$?
- 2. Why does variable density sampling outperform uniform random sampling?

Remark: Feasible sampling patterns



In practical applications such as MRI, the hardware constraints mean that uniform random sampling cannot be implemented. Thus, there is a need to understand how one should sample along trajectories.

Last part of this talk: Sampling along Cartesian lines.

Notation

1D case: For $x \in \mathbb{C}^N$, $\Lambda \subset \mathbb{Z}$, • $Ux = \left(\sum_{j=1}^N x_j e^{i2\pi k_j}\right)_{k=-\lfloor N/2 \rfloor, \dots, \lceil N/2 \rceil - 1}$. • $P_\Lambda : \mathbb{C}^N \to \mathbb{C}^N$, $(P_\Lambda x)_j = x_j$ if $j \in \Lambda$ and 0 otherwise. • $Dx := (-x_j + x_{j+1})_{j=1}^N$ with $x_{N+1} := x_1$. Let $\|x\|_{TV} = \|Dx\|_1$.

2D case: For $x \in \mathbb{C}^{N \times N}$, $\Lambda \subset \mathbb{Z}^2$,

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$$Ux = \left(\sum_{j_1=1}^{N} \sum_{j_2=1}^{N} x_{j_1,j_2} e^{i2\pi(k_1j_1+k_2j_2)}\right)_{k_1,k_2=-\lfloor N/2 \rfloor,\dots,\lceil N/2 \rceil-1}$$

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with $x_{N+1,j} := x_{1,j}$ and $x_{k,N+1} := x_{k,1}$. Let $||x||_{TV} := ||Dx||_1$.

Given y and $\Omega \subset \mathbb{Z}$ (resp. \mathbb{Z}^2) of cardinality m, we will consider the solutions $\mathcal{R}(\Omega, \delta, y) = \operatorname*{argmin}_{z \in \mathbb{C}^N (\operatorname{resp.} \mathbb{C}^{N \times N})} \|z\|_{TV}$ subject to $\|P_{\Omega}U - y\|_2 \leq \sqrt{m} \cdot \delta$.

Outline

Uniform + power law sampling

Uniform random sampling

Beyond sparsity - low frequency sampling

Sampling along Cartesian lines

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Definition

 $\Omega = \Omega_1 \cup \Omega_2 \subset \{-\lfloor N/2 \rfloor + 1, \dots, \lceil N/2 \rceil\}$ is a uniform + power law sampling scheme of cardinality 2m if

- Ω_1 consists of *m* indices chosen uniformly at random.
- $\Omega_2 = \{k_1, \ldots, k_m\}$ consist of *m* indices which are independent and identically distributed (i.i.d.) such that for each $j = 1, \ldots, m$ and $n = -N/2 + 1, \ldots, N/2$,

$$\mathbb{P}(k_j = n) = p(n), \quad p(n) = \frac{C \log(N)}{\max\{1, |n|\}},$$

where C is an appropriate constant such that p is a probability measure.



Theorem (P. 2015)

Let $N = 2^J$ with $J \in \mathbb{N}$, $\delta \ge 0$ and $\epsilon \in (0, 1)$.

- Let $x \in \mathbb{C}^N$ and let Δ index the largest s coefficients of Dx.
- ► Suppose we are given $y = P_{\Omega}Ux + \eta$ where $\|\eta\|_2 \le \sqrt{m} \cdot \delta$ and Ω is a uniform + power law sampling scheme of cardinality $m = \mathcal{O}\left(s \log(N)(1 + \log(\epsilon^{-1}))\right)$.

Then with probability exceeding $1 - \epsilon, \xi \in \mathcal{R}(\Omega, \delta, y)$ satisfies

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where $\mathcal{L}_1 = \log^2(s) \log(N) \log(m)$ and $\mathcal{L}_2 = \log(s) \log^{\frac{1}{2}}(m)$.

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$$\begin{split} \|Dx - D\xi\|_2 \lesssim \left(\delta\sqrt{s} + \mathcal{L}_2 \cdot \frac{\|P_{\Delta^c} Dx\|_1}{\sqrt{s}}\right), \\ \frac{\|x - \xi\|_2}{\sqrt{N}} \lesssim \mathcal{L}_1 \cdot \left(\frac{\delta}{\sqrt{s}} + \mathcal{L}_2 \cdot \frac{\|P_{\Delta^c} Dx\|_1}{s}\right), \end{split}$$

where $\mathcal{L}_1 = \log^2(s) \log(N) \log(m)$ and $\mathcal{L}_2 = \log(s) \log^{\frac{1}{2}}(m)$.

DeVore (1998): The optimal error decay rate for any bounded variation function $f \in BV[0, 1)$ by any type of nonlinear approximation \tilde{f} from s samples is

$$\|\tilde{f} - f\|_{L^2[0,1)} = \mathcal{O}\left(\|f\|_V \cdot s^{-1}\right).$$

Definition

 $\Omega = \Omega_1 \cup \Omega_2 \subset \{-\lfloor N/2 \rfloor + 1, \dots, \lceil N/2 \rceil\}^2$ is a uniform + power law sampling scheme of cardinality 2m if

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$$\mathbb{P}(k_j = (n_1, n_2)) = p(n_1, n_2),$$
$$p(n_1, n_2) = \frac{C \log(N)}{\max\{1, |n_1|^2 + |n_2|^2\}}$$

where C > 0 is such that p is a probability measure.

Theorem (P. 2015) Let $N = 2^J$, $J \in \mathbb{N}$, $\epsilon \in (0, 1)$ and $\delta \ge 0$.

- Let $x \in \mathbb{C}^{N \times N}$ and let Δ index the largest s coefficients of Dx.
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$$m = \mathcal{O}\left(s\log(N)(1 + \log(\epsilon^{-1}))\right).$$

Then, with probability exceeding $1 - \epsilon, \xi \in \mathcal{R}(\Omega, \delta, y)$ satisfies

$$\begin{split} \|Dx - D\xi\|_2 \lesssim \left(\delta \cdot \sqrt{s} + \mathcal{L}_2 \cdot \frac{\|P_{\Delta^c} Dx\|_1}{\sqrt{s}}\right), \\ \|x - \xi\|_2 \lesssim \mathcal{L}_1 \cdot \left(\delta + \mathcal{L}_2 \cdot \frac{\|P_{\Delta^c} Dx\|_1}{\sqrt{s}}\right), \\ \end{split}$$
 where $\mathcal{L}_1 = \log(s) \log(\frac{N^2}{s}) \log^{\frac{1}{2}}(N) \log^{\frac{1}{2}}(m)$, and $\mathcal{L}_2 = \log^{\frac{1}{2}}(m) \log(s).$

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$$\|Dx - D\xi\|_{2} \lesssim \left(\delta \cdot \sqrt{s} + \mathcal{L}_{2} \cdot \frac{\|P_{\Delta^{c}} Dx\|_{1}}{\sqrt{s}}\right),$$
$$\|x - \xi\|_{2} \lesssim \mathcal{L}_{1} \cdot \left(\delta + \mathcal{L}_{2} \cdot \frac{\|P_{\Delta^{c}} Dx\|_{1}}{\sqrt{s}}\right),$$
$$(c) \log^{(N^{2})} \log^{\frac{1}{2}}(N) \log^{\frac{1}{2}}(m) \text{ and } \mathcal{L}_{1} = \log^{\frac{1}{2}}(m)$$

where $\mathcal{L}_1 = \log(s) \log(\frac{N^2}{s}) \log^{\frac{1}{2}}(N) \log^{\frac{1}{2}}(m)$, and $\mathcal{L}_2 = \log^{\frac{1}{2}}(m) \log(s)$.

Candès & Tao (2006), Needell & Ward (2013): The optimal error estimate from $O(s \log(N^2/s))$ nonadaptive samples is

$$\delta + \frac{\|P_{\Delta^c} Dx\|_1}{\sqrt{s}}.$$

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Krahmer & Ward (2014): Given $\mathcal{O}(s \log^5(N) \log^3(s))$ Fourier coefficients distributed by a power law, TV regularization guarantees stable recovery up to gradient sparsity s.

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Uniform random sampling (1D case)

Theorem (P. 2015)

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- ► Suppose we are given $y = P_{\Omega}Ux + \eta \in \mathbb{C}^m$ where $\|\eta\|_2 \leq \sqrt{m} \cdot \delta$ and Ω includes 0 and *m* indices chosen uniformly at random with

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where $\mathcal{L} = \log^{\frac{1}{2}}(m) \log(s)$.

Uniform random sampling (2D case)

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Question...

So, uniform random sampling does achieve robustness and stability...

Example: Reconstructions from 35% uniform random sampling:



However, the error bounds obtained for the uniform random sampling strategy are sub-optimal, whereas, by adding the samples which concentrate on low frequencies, one can guarantee near-optimal error bounds.

Does dense sampling at low frequencies actually improve stability, or is the difference between the theorems simply an artefact of the proofs?

A numerical comparison

Consider the recovery of x+h from 10% of its Fourier coefficients, with different $SNR=10\log_{10}(\|x\|_2\,/\,\|h\|_2).$



Relative errors of the recovered signals using Ω_P vs using Ω_U .

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Detour: super resolution (discrete case)

The recovery of a super position of spikes, $x = \sum_{j} \alpha_{j} \delta_{t_{j}}$ with $t_{j} \in [0, 1]$, from low frequency samples only.

Candès & Fernandez-Granda (2012): Let $x \in \mathbb{C}^N$, let Δ be its support and suppose that

$$\min_{t,t'\in\Delta,t\neq t'}\frac{|t-t'|}{N}\geq \frac{2}{M}$$

Then, given $y = P_{[M]}Ux$ (first 2M + 1 DFT coefficients of x) with $[M] = \{-M, \ldots, M\}$, x is the unique solution of

 $\min_{z} \|z\|_1 \text{ subject to } P_{[M]}Uz = y.$

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If instead, Dx has support Δ :

- 1. recover Dx by solving an ℓ^1 problem.
- 2. recover x from Dx by shifting by the mean of x (i.e. 0^{th} Fourier coefficient of x).

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Tang & Bhaskar & Shah & Recht (2013): An analogous compressed sensing result.

Low frequency sampling

Theorem (P. 2015)

Let $N \in \mathbb{N}$, let $\epsilon \in [0, 1]$ and let $M \in \mathbb{N}$ be such that $N/4 \ge M \ge 10$.

• Let $x \in \mathbb{C}^N$ and $\Delta \subset \{1, \dots, N\}$ be of cardinality s and suppose that

$$\min_{k,j\in\Delta,k\neq j}\frac{|k-j|}{N}\geq \frac{2}{M}$$

 \blacktriangleright Let $\Omega \subset \{-M, \ldots, M\}$ include 0 and m indices chosen uniformly at random with

$$m \gtrsim \max\left\{\log^2\left(\frac{M}{\epsilon}\right), \quad \log(N), \quad s \cdot \log\left(\frac{s}{\epsilon}\right) \cdot \log\left(\frac{M}{\epsilon}\right)\right\}.$$

Then with probability exceeding $1 - \epsilon$, given $y = P_{\Omega}Ax + \eta$ and $\|\eta\|_2 \leq \delta \cdot \sqrt{m}$, any solution $\xi \in \mathcal{R}(\Omega, \delta, y)$ satisfies

$$\frac{\|x - \xi\|_2}{\sqrt{N}} \lesssim \frac{N^2}{M^2} \cdot \left(\delta \cdot s + \sqrt{s} \cdot \|P_{\Delta^c} Dx\|_1\right).$$

If m = 2M + 1, then the error bound holds with probability 1.

The price of randomness

Suppose that $x \in \mathbb{C}^N$ was s-gradient sparse with a minimum separation of 2/s. Then, x can be exactly recovered from 2s + 1 Fourier coefficients. However, random sampling guarantees recovery only with $\mathcal{O}(s \log N)$ samples.



This signal can be recovered exactly from 3.9% of its Fourier coefficient of lowest frequencies, but uniform random sampling would require sampling at 10%.

What about the 2D case?

Candès & Fernandez-Granda (2012):

Let $x \in \mathbb{C}^{N \times N}$ have support Δ . If $\min_{\mathbf{k}, \mathbf{j} \in \Delta, \mathbf{k} \neq \mathbf{j}} |\mathbf{k} - \mathbf{j}| \geq 2.38/M$ and we observe the Fourier coefficients of x up to frequency $M \in \mathbb{N}$, $y = P_{[M]}Ux$. Then x is the unique solution of

 $\min_{z} \|z\|_1 \text{ subject to } P_{[M]}Uz = y.$



The difficulty with the recovery of images is that there is no separation in the edge set.

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In 1D, the *number* of samples depends on the *gradient sparsity*, and the *range* that we sample from depends on the *separation* of the support set. In 2D:

- ▶ The number of lines depends on the gradient sparsity along each direction.
- ▶ The sampling range depends on the separation along each direction.



Separation concepts

Definition

Let $N \in \mathbb{N}$ and let $\Delta \subset \{1, \dots, N\}^2$. The minimum separation distance of

▶ its rows is defined to be

$$\nu_{\text{row}}(\Delta, N) = \min_{n=1}^{N} \min\left\{\frac{|j-k|}{N} : j, k \in \Delta \cap \{\{n\} \times \{1, \dots, N\}\}\right\},\$$

▶ its columns is defined to be

$$\nu_{\rm col}(\Delta,N) = \min_{n=1}^N \min\left\{\frac{|j-k|}{N} : j,k \in \Delta \cap \{\{1,\ldots,N\} \times \{n\}\}\right\}.$$



Sparsity concepts

Definition Let $\Delta \subset \{1, \dots, N\}^2$. Δ is of cardinality s

▶ along its columns if $s = \max_{j=1}^{N} \left| \Delta_{j}^{[\text{col}]} \right|$, where

 $\Delta_j^{[\text{col}]} = \{ (n_1, j) \in \Delta : n_1 = 1, \dots, N \} \,.$

► along its rows if $s = \max_{j=1}^{N} \left| \Delta_{j}^{[\text{row}]} \right|$, where

$$\Delta_j^{[\text{row}]} = \{(j, n_2) \in \Delta : n_2 = 1, \dots, N\}.$$

Definition

Let $x \in \mathbb{C}^{N \times N}$. Let $x^{[\operatorname{col},j]}$ be the j^{th} column of x and $x^{[\operatorname{row},j]}$ be the j^{th} row of x. We say that x has T distinct supports

▶ along its columns if $T = \left| \left\{ x^{[\text{col},j]} : j = 1, \dots, N \right\} \right|.$

• along its rows if
$$T = \left| \left\{ x^{[row,j]} : j = 1, \dots, N \right\} \right|.$$

Our assumptions

Let $x \in \mathbb{C}^{N \times N}$ and let $\Delta_1, \Delta_2 \subset \{1, \ldots, N\}^2$ index the largest s_1 coefficients of $D_1 x$ and the largest s_2 coefficients of $D_2 x$ resp.

- Along its columns, Δ_1 has a minimum separation of $2/M_1$ and is of cardinality s_1 .
- Along its rows, Δ_2 has a minimum separation of $2/M_2$ and is of cardinality s_2 .
- ▶ P_{Δ_1} sgn (D_1x) has T_1 distinct supports along its columns.
- ▶ P_{Δ_2} sgn $(D_2 x)$ has T_2 distinct supports along its rows.

For a Cartesian line sampling index set Ω with $|\Omega| = m$ and $y = P_{\Omega}x + \eta$ with $||\eta|| \le \delta\sqrt{m}$, we will consider the solutions of

$$\min_{z \in \mathbb{C}^{N \times N}} \|z\|_{TV,\text{aniso}} \text{ subject to } \|P_{\Omega}Uz - y\| \le \delta \sqrt{m}$$

Recall that D_1 performs finite differences along each column and D_2 performs finite differences along each row. The anisotropic total variation norm of $x \in \mathbb{C}^{N \times N}$ is

$$\|x\|_{TV,\text{aniso}} := \|D_1 x\|_1 + \|D_2 x\|_1 = \sum_{\mathbf{j}} |(D_1 x)_{\mathbf{j}}| + |(D_2 x)_{\mathbf{j}}|,$$

as opposed to the isotropic total variation norm $||x||_{TV} = \sum_{\mathbf{j}} \sqrt{|(D_1x)_{\mathbf{j}}|^2 + |(D_2x)_{\mathbf{j}}|^2}$.

Sampling along Cartesian lines

Theorem (P. 2015) Let $\epsilon \in (0, 1)$. Let $\Omega = \{0\} \cup \{\Omega_1 \times [N]\} \cup \{[N] \times \Omega_2\}$, and $m = |\Omega|$, where $\Omega_1 \sim \text{Unif}([M_1], m_1), \qquad \Omega_2 \sim \text{Unif}([M_2], m_2),$ $m_1 \gtrsim \max\left\{\log^2(T_1M_1/\epsilon), \quad \log(N), \quad s_1\log(T_1s_1/\epsilon)\log(T_1M_1/\epsilon)\right\},$ and

 $m_2 \gtrsim \max\left\{\log^2(T_2M_2/\epsilon), \quad \log(N), \quad s_2\log(T_2s_2/\epsilon)\log(T_2M_2/\epsilon)\right\}.$

Then, with probability exceeding $1 - \epsilon$, any minimizer \hat{x} satisfies

$$\|D(x-\hat{x})\|_{2} \lesssim \frac{N^{2}}{M_{0}^{2}} \left((m_{0}N)^{-1/2} \sqrt{m}\delta + \left\| P_{\Delta_{1}}^{\perp} D_{1}x \right\|_{1} + \left\| P_{\Delta_{2}}^{\perp} D_{2}x \right\|_{1} \right),$$

and

$$\|x - \hat{x}\| \lesssim \frac{N^2}{M_0^2} \sqrt{s} \left((m/m_0)^{1/2} \delta + \left\| P_{\Delta_1}^{\perp} D_1 x \right\|_1 + \left\| P_{\Delta_2}^{\perp} D_2 x \right\|_1 \right),$$

where $s = \max\{s_1, s_2\}, m_0 = \min\{m_1, m_2\}$, and $M_0 = \min\{M_1, M_2\}$.

If $\Omega_1 = [M_1]$ and $\Omega_2 = [M_2]$, then these bounds hold with probability one.

Example: Sampling 1.2% of the Fourier coefficients

If x has at most s_1 discontinuities along each of its columns with a minimum separation of $2/s_1$ and it has at most s_2 discontinuities along each of its rows with a minimum separation of $2/s_2$, then one is guaranteed exact recovery by sampling along $2(s_1 + s_2)$ Cartesian lines.



Sampling in accordance to sparsity structure allows for $\operatorname{sub-}\mathcal{O}(s \log N)$ recovery.

Reconstruction of the 1951 USAF resolution test chart (6.5% sampling)



Conclusions

Although uniform random sampling is stable and robust...

- ▶ a uniform + power law sampling strategy achieves recovery guarantees which are optimal (for sparse vectors) up to log factors.
- ▶ in the 1D case where the discontinuities of the underlying signal are sufficiently far apart, one only needs to sample from low Fourier frequencies to ensure exact recovery.
- ▶ in the 2D case, recovery guarantees were presented for sampling along Cartesian lines. The sampling result is dependent both on sparsity and the sparsity separation in each direction.
- by accounting for sparsity structure, one can circumvent the $\mathcal{O}(s \log N)$ bound.
- ▶ variable density sampling schemes appear to combine the benefits of super resolution and compressed sensing: allows for a linear correspondence between the coarse features recovered and the number of samples, and also the recovery of fine features at the price of a log factor.

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Thanks for listening!

On the role of total variation in compressed sensing. SIAM J. Imaging Sci., 8(1), 682-720, 2015.

Remark on the proofs

One can show that if $x \in \mathbb{C}^N$ and $\operatorname{supp}(Dx) = \Delta$, then x is the unique solution to

$$\min_{z} \|z\|_{TV} \text{ subject to } P_{\Omega}Uz = P_{\Omega}Uz,$$

provided that

- 1. $P_{\Omega}UP_{\Delta}$ is injective,
- 2. There exists $\eta \in \operatorname{ran}(U^*P_\Omega)$ such that $\eta_j = \operatorname{sgn}(Dx)_j$ for all $j \in \Delta$ and $\|\eta\|_{\infty} \leq 1$.
- 3. $\|h\| \leq C(N) \|h\|_{TV}$ whenever $P_{\Omega}Uh = 0$.
- ▶ The second condition is simply asking if there exists a trigonometric polynomial

$$p = \sum_{j \in \Omega} \alpha_j e^{2\pi i \langle j, \cdot \rangle}$$

which interpolates the sign pattern of Dx and $||p||_{\infty} \leq 1$.

• The third condition is true if $0 \in \Omega$, and one can show that this condition holds with smaller constants C(N) under power law sampling.